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## A MODIFIED MODEL OF RISK BUSINESS

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We consider the risk model in which the claim counting process  $\{N(t)\}$  is a modified stationary renewal process.  $\{N(t)\}$  is governed by a sequence of independent and identically distributed inter-occurrence times with a common exponential distribution function with mass at zero equal to  $\rho > 0$ . The model is called a Pólya - Aepli risk model. The Cramér - Lundberg approximation and the martingale approach of the model are given.

### 1. Introduction

Assume that the standard model of an insurance company, called risk process  $\{X(t), t \geq 0\}$  is given by

$$(1) \quad X(t) = ct - \sum_{k=1}^{N(t)} Z_k, \quad \left( \sum_1^0 = 0 \right).$$

Here  $c$  is a positive real constant representing the risk premium rate. The sequence  $\{Z_k\}_{k=1}^{\infty}$  of mutually independent and identically distributed random variables with common distribution function  $F$ ,  $F(0) = 0$ , and finite mean value  $\mu$  is independent of the counting process  $N(t)$ ,  $t \geq 0$ . The process  $N(t)$  is interpreted as the number of claims on the company during the interval  $[0, t]$ . In

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the classical risk model the process  $N(t)$  is a homogeneous Poisson process, see for instance [1] and [6].

The Pólya - Aeppli distribution is a generalization of the classical  $Po(\lambda)$  distribution, by adding a new parameter  $\rho$ , see [2]. It appears in [4] and [5] as a compound Poisson distribution. The additional parameter  $\rho$  is called an inflation parameter. The Pólya - Aeppli process as a generalization of the Poisson process is defined in [3].

We will suppose that  $N(t)$  is described by the Pólya - Aeppli distribution with mean function  $\frac{\lambda}{1-\rho}t$ , i.e.

$$(2) \quad P(N(t) = n) = \begin{cases} e^{-\lambda t}, & n = 0 \\ e^{-\lambda t} \sum_{i=1}^n \binom{n-1}{i-1} \frac{[\lambda(1-\rho)t]^i}{i!} \rho^{n-i}, & n = 1, 2, \dots \end{cases}$$

In this section we will discuss briefly the basic properties of the Pólya - Aeppli process.

In [3] is proved that the non-negative random variables  $T_1, T_2, \dots$ , representing the inter-arrival times are mutually independent. The time  $T_1$  until the first epoch is exponentially distributed with parameter  $\lambda$ .  $T_2, T_3, \dots$  are identically distributed as a random variable  $T_2$ . Moreover  $T_2$  is zero with probability  $\rho$ , and with probability  $1 - \rho$  exponentially distributed with parameter  $\lambda$ , i.e.  $T_2$  is exponentially distributed with mass at zero equal to  $\rho$ . The probability distribution function is given by

$$F_{T_2}(t) = 1 - (1 - \rho)e^{-\lambda t}, \quad t \geq 0.$$

The mean value is  $ET_2 = \frac{1-\rho}{\lambda}$ .

The process, described above is a delayed renewal process. It is easy to verify that the probability distribution functions of the delay  $T_1$  and the inter-arrival times satisfy the following relation

$$F_{T_1} = \frac{1}{ET_2} \int_0^t [1 - F_{T_2}(u)] du.$$

In this case the delayed renewal process is the only stationary renewal process. So, the Pólya - Aeppli process is a homogeneous process and if  $\rho = 0$  it becomes a homogeneous Poisson process.

In this paper we need also the probability generating function (pgf) of the Pólya - Aeppli process. It is given by

$$P_{N(t)}(s) = e^{-\lambda t \frac{1-s}{1-\rho s}}.$$

The Pólya - Aeppli risk model is defined in [3]. The probability of ruin and the Cramér - Lundberg approximation are derived. In this paper the martingale approach is given.

## 2. The Pólya - Aeppli risk model

We consider the risk process  $X(t)$ , defined by (1), where  $N(t)$  is the Pólya - Aeppli process, independent of the claim sizes  $Z_k, k = 1, 2, \dots$ . This process is called a Pólya - Aeppli risk model.

The relative safety loading  $\theta$  is defined by

$$(3) \quad \theta = \frac{c(1 - \rho) - \lambda\mu}{\lambda\mu} = \frac{c(1 - \rho)}{\lambda\mu} - 1,$$

and in the case of positive safety loading  $\theta > 0, c > \frac{\lambda\mu}{1-\rho}$ .

Let

$$F_I(x) = \frac{1}{\mu} \int_0^x [1 - F(z)] dz$$

be the integrated tail distribution of the claim size. Let us define the function

$$(4) \quad H(z) = \rho F(z) + \frac{\lambda\mu}{c} F_I(z)$$

and note that

$$H(\infty) = \rho F(\infty) + \frac{\lambda}{c} \int_0^\infty [1 - F(z)] dz = \rho + \frac{\lambda\mu}{c} < 1.$$

Denote by

$$\tau(u) = \inf\{t > 0, u + X(t) \leq 0\}$$

the time to ruin of a company having initial capital  $u$ . We let  $\tau = \infty$ , if for all  $t > 0, u + X(t) > 0$ .

The probability of ruin in the infinite horizon case is

$$\Psi(u) = P(\tau(u) < \infty)$$

and in the finite horizon case

$$\Psi(u, t) = P(\tau(u) \leq t).$$

Assume that there exists a constant  $R > 0$  such that

$$(5) \quad \int_0^\infty e^{Rz} dH(z) = 1,$$

where  $H(z)$  is given by (4), and denote  $h(R) = \int_0^\infty e^{Rz} dF(z) - 1$ .

The following proposition, proved in [3] is an analogue of the corresponding approach to the classical risk model.

**Proposition 1.** *Let, for the Pólya-Aeppli risk model, the Cramér condition (5) holds and  $h'(R) < \infty$ . Then*

$$(6) \quad \lim_{u \rightarrow \infty} \Psi(u) e^{Ru} = \frac{\mu\theta}{A^2(\mu, \theta, R, \rho)h'(R) - \mu(1 + \theta)},$$

where

$$A(\mu, \theta, R, \rho) = \frac{1 - [1 - \mu(1 + \theta)R]\rho}{1 - \rho}.$$

The relation (5) is known as the Cramér condition and the constant  $R$ , if it exists, as the adjustment coefficient or Lundberg exponent for the Pólya-Aeppli risk model.

If  $\rho = 0$ ,  $A(\mu, \theta, R, 0) = 1$  and (6) coincides with the Cramér - Lundberg approximation for the classical risk model [1].

### 3. Martingales for the Pólya-Aeppli risk model

Let us denote by  $(\mathcal{F}_t^X)$  the natural filtration generated by any stochastic process  $X(t)$ .  $(\mathcal{F}_t^X)$  is the smallest complete filtration to which  $X(t)$  is adapted.

Let us denote by  $LS_Z(r) = \int_0^\infty e^{-rx} dF(x)$  the Laplace-Stieltjes transform (LS-transform) of any random variable  $Z$  with distribution function  $F(x)$ .

**Lemma 1.** *For the Pólya-Aeppli risk model*

$$Ee^{-rX(t)} = e^{g(r)t},$$

where

$$g(r) = \frac{1}{1 - \rho LS_Z(-r)} [\rho cr LS_Z(-r) + \lambda (LS_Z(-r) - 1) - cr].$$

**Proof.** Let us consider the random sum from the right hand side of (1)

$$S_t = \sum_{k=1}^{N(t)} Z_k,$$

where  $N(t)$  is a Pólya-Aeppli process, independent of  $Z_k, k = 1, 2, \dots$ .  $S_t$  is a compound Pólya-Aeppli process and the LS- transform is given by

$$LS_{S_t}(r) = P_{N(t)}(LS_Z(r)) = e^{-\lambda t \frac{1-LS(r)}{1-\rho LS(r)}}.$$

For the LS-transform of  $X(t)$  we have the following

$$\begin{aligned} LS_{X(t)}(r) &= Ee^{-rX(t)} = Ee^{-r[ct-S_t]} = e^{-rct} Ee^{rS_t} = \\ &= e^{-rct} P_{N(t)}(LS_Z(-r)) = e^{-rct} e^{-\lambda t \frac{1-LS_Z(-r)}{1-\rho LS_Z(-r)}} = e^{g(r)t}. \end{aligned}$$

□

From the martingale theory we get the following

**Lemma 2.** For all  $r \in R$  the process

$$M_t = e^{-rX(t)-g(r)t}, \quad t \geq 0$$

is an  $\mathcal{F}_t^X$ -martingale, provided that  $LS_Z(-r) < \infty$ .

#### 4. Martingale approach to the Pólya-Aeppli risk model

Using the martingale properties of  $M_t$ , we will give some useful inequalities for the ruin probability.

**Proposition 2.** Let  $r > 0$ . For the ruin probabilities of the Pólya-Aeppli risk model we have the following results

i)  $\Psi(u, t) \leq e^{-ru} \sup_{0 \leq s \leq t} e^{g(r)s}, 0 \leq t < \infty$

ii)  $\Psi(u) \leq e^{-ru} \sup_{s \geq 0} e^{g(r)s}.$

iii) If the Lundberg exponent  $R$  exists, then  $R$  is the unique strictly positive solution of

$$(7) \quad \rho cr LS_Z(-r) + \lambda (LS_Z(-r) - 1) - cr = 0$$

and  $\Psi(u) \leq e^{-Ru}.$

Proof. i) For any  $t_0 < \infty$ , the martingale stopping time theorem yields the following

$$\begin{aligned} 1 = M_0 &= EM_{t_0 \wedge \tau} = E[M_{t_0 \wedge \tau}, \tau \leq t] + E[M_{t_0 \wedge \tau}, \tau > t] \geq \\ &\geq E[M_{t_0 \wedge \tau}, \tau \leq t] = E[e^{-rX(\tau)-g(r)\tau} | \tau \leq t] P(\tau \leq t), \end{aligned}$$

from which

$$P(\tau \leq t) = \frac{e^{-ru}}{E[e^{-g(r)\tau} | \tau \leq t]}.$$

The statement i) follows from the above relation.

ii) follows immediately from i) when  $t \rightarrow \infty$ .

iii) The Cramér condition (5) becomes

$$\rho \int_0^\infty e^{rx} dF(x) + \frac{\lambda}{c} \int_0^\infty e^{rx} (1 - F(x)) dx = 1.$$

Using

$$\int_0^\infty e^{rx} (1 - F(x)) dx = \int_0^\infty \int_x^\infty e^{rx} dF(y) dx = \int_0^\infty \int_0^y e^{rx} dx dF(y) = \frac{1}{r} [LS_Z(-r) - 1],$$

it can be written as

$$\rho LS_Z(-r) + \frac{\lambda}{cr} (LS_Z(-r) - 1) = 1.$$

This is equivalent to the equation (7).

Let us denote

$$f(r) = \rho cr LS_Z(-r) + \lambda (LS_Z(-r) - 1) - cr.$$

So  $R$  is a positive solution of  $f(r) = 0$ . Because  $f(0) = 0$ ,  $f'(0) = \lambda\mu - (1 - \rho)c < 0$  and  $f''(r) = (\rho cr + \lambda)LS_Z''(-r) + 2\rho c LS_Z'(-r) > 0$  there is at most one strictly positive solution.  $\square$

**Remark 1.** The equation (7) is equivalent to  $g(r) = 0$ .

**Remark 2.** The above inequalities are well known for the classical risk model. In the case of  $\rho = 0$  the Pólya - Aepli risk model becomes the classical risk model. The Cramér condition (5) and the function  $g(r)$  are the same, see [7].

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