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CONTROLLED PROCESS WITH RANDOM BREAKDOWNS AND REPEAT ACTIONS

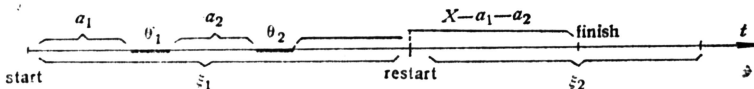
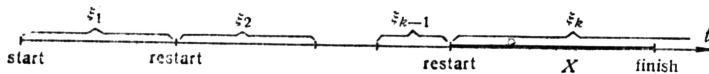
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Some models minimizing the duration of the unreliable process by the help of statement copies are considered.

1. Introduction. Let during the operating time of a process (for example by computing, by constructing) in a random way some undesirable events (breakdowns, catastrophes) arise, which leads to process interruption. If one wants to have a successful finish, the process must be repeated from the origin (Fig. 1). It is profitable to introduce a strategy for making copies and remember the process states in some chosen moments. If a catastrophe appear the process continues from the last copied state. In the present paper a strategy for process control is proposed. This strategy minimizes a functional of the pure process operating time.

2. The basic model. Let X be the pure duration of the process without breakdowns, $\{\xi_n\}$ a random flow of discrete events (catastrophes) and $\tau(X)$ the full duration of the process up to finish. The process starts at $t=0$. If the process restarts after every catastrophe and k is the first number k for which $\xi_k \geq X$, then $\tau(X) = \sum_{i=1}^{k-1} \xi_i + X$. Evidently, $\tau(X) \geq X$.

3. The copies control. Let us consider a sequence $\{a_k\}$ of time intervals between copies and the sequences $\{\theta_k\}$ of copying time durations. After every time interval a_i one makes a copy of the process state in a time θ_i . The breakdowns are possible and during the copies. The process restarts from the last copying state (Fig. 2). In that case the full process duration is denoted by $\tau(X, \{a_k\}, \{\theta_k\})$ (controlled time duration). Sometimes the short notation $\tau(X)$ will be used (e. g. in the expectations).



4. Some simplifying suppositions. Under general assumptions no one can define any problem. We suppose, that X is a random variable (r.v.) with a given distribution function (d.f.) $A(x) = P\{X < x\}$; $\{\xi_n\}$ is a Poisson process with an intensity $\gamma > 0$; $\{a_k\}$ is a deterministic process, $\{\theta_k\}$ forms a renewable one, and $P\{\theta_k < x\} = F(x)$. Then $\tau(X, \{a_k\}, \{\theta_k\}) = \tau(X, \{a_k\}, \lambda, \theta)$ is a r.v.

5. The optimal control problem. Under the above assumptions the question is: how to choose $\{a_k^*\}$ so that

$$\inf_{\{a_k\}} E \tau(X, \{a_k\}, \theta) = E \tau(X, a_k^*, \theta).$$

The sequence $\{a_k^*\}$ will be called *optimal copy schedule*.

6. Some exact results. Further we suppose the assumptions 5 are fulfilled. We introduce the functions

$$\tau^*(s) = E e^{-s\tau(X)}, \quad \tau^{**}(s) = E \exp[-s\tau(X, \{a_k\}, \theta)],$$

which are the Laplace-Stieltjes (L.S.) transformations of the probable distributions of r.v. $\tau(X)$ and $\tau(X, \{a_k\}, \theta)$ accordingly.

6.1. The case $X = \text{const.}$

Lemma 1. *If $P\{X = x\} = 1$, then*

$$\tau^*(s) = \frac{s + \gamma}{\gamma + s \exp[(s + \gamma)x]}.$$

The proof is a simple consequence of the next recurrent relation:

$$(1) \quad \tau^*(s) = e^{-(s + \gamma)x} + \frac{\gamma}{s + \gamma} [1 - e^{-(s + \gamma)x}] \tau^*(s).$$

One can derive (1), using the probabilistic interpretation of L.S. transformations [5].

Lemma 2. *If $P\{X = x\} = 1$, then*

$$(2) \quad \tau^{**}(s) = \frac{s + \gamma}{\gamma + s \cdot \exp[(s + \gamma)(x - \sum_{k=1}^{n_x} a_k)]} \prod_{k=1}^{n_x} \frac{(s + \gamma)\varphi(s + \gamma)}{\gamma\varphi(s + \gamma) + s \cdot \exp[a_k(s + \gamma)]},$$

where $\varphi(s) = E\{\exp(-s\theta_k)\}$, $n_x = \max[n; a_1 + a_2 + \dots + a_n \leq x]$.

Proof. First of all, we mention, that a copy interval a_k and the next copy time θ_k form an interval of length τ_k in the full process duration $\tau(X, \{a_k\}, \theta)$. There are n_x such of intervals and the last one doesn't need a copy time; its duration is equal to $\tau(x - \sum_{k=1}^{n_x} a_k)$. The "lack of memory" property of the exponential distribution shows, that the r.v. $\{\tau_k\}$ are independent, and

$$(3) \quad \tau(x, \{a_k\}, \theta) = \sum_{k=1}^{n_x} \tau_k + \tau(x - \sum_{k=1}^{n_x} a_k).$$

Further, we introduce $\tau_k^*(s) = E e^{-s\tau_k}$. Using the "lack of memory" property, we get again

$$\tau_k^*(s) = e^{-(s + \gamma)a_k} \varphi(s + \gamma) + \frac{\gamma}{s + \gamma} [1 - e^{-(s + \gamma)a_k} \varphi(s + \gamma)] \tau_k^*(s).$$

It gives the relation

$$(4) \quad \tau_k^*(s) = \frac{(s+\gamma)\varphi(s+\gamma)}{\gamma\varphi(s+\gamma) + se^{(s+\gamma)a_k}}.$$

Now (2) follows from Lemma 1, (3) and (4).

Lemma 3. *The expected full time duration of the copy controlled process in the case $P\{X=x\}=1$ is*

$$E\tau(x, \{a_k\}, \theta) = \frac{1}{\gamma} \sum_{k=1}^{n_x} \left[\frac{e^{\gamma a_k}}{\varphi(\gamma)} - 1 \right] + \frac{1}{\gamma} \left[e^{\gamma(x - \sum_{k=1}^{n_x} a_k)} - 1 \right].$$

The proof follows from Lemma 2 and the well-known relation

$$E\xi = -\frac{d}{dx} (E e^{-s\xi})|_{s=0}.$$

Theorem 1. *If $P\{X=x\}=1$ then the optimal copy schedule satisfies the condition $a_k^* = a^* = \text{const}$.*

Proof. Let us denote $b_0 = 0$, $b_k = a_1 + a_2 + \dots + a_k$ and $n_x = \max(k, b_k \leq x)$. Then $b_k + k\theta$ is the moment of k -th copy from the origin, if there is no catastrophes. The result of Lemma 3 can be rewritten in the form

$$(5) \quad E\tau(x, \{b_k\}, \theta) = \frac{1}{\gamma} \sum_{k=1}^{n_x} \left[\frac{e^{\gamma(b_k - b_{k-1})}}{\varphi(\gamma)} - 1 \right] + \frac{1}{\gamma} \left[e^{\gamma(x - b_{n_x})} - 1 \right].$$

The optimal copy schedule $\{a_k^*\}$ gives the sequence $\{b_k^*\}$, which is a solution of the system of equations

$$(6) \quad \frac{\partial E\tau(x, \{a_k\}, \theta)}{\partial b_k} = 0, \quad k = 1, 2, \dots$$

It is easy to see, that in the case (5) the system (6) becomes the form $\frac{1}{\varphi(\gamma)} e^{\gamma(b_k - b_{k-1})} - \frac{1}{\varphi(\gamma)} e^{\gamma(b_{k+1} - b_k)} = 0$, $k = 1, 2, \dots$, which proves the theorem. For an abbreviation we introduce the notations $k_x = k(a, x) = [x/a]$ the integer part of x/a ;

$$I_k(a) = \begin{cases} 1 & \text{if } a \in [x/(k+1), x/k], \\ 0 & \text{otherwise;} \end{cases}$$

$$T_k(x, a) = \frac{k}{\gamma} \left[\frac{e^{\gamma a}}{\varphi(\gamma)} - 1 \right] + \frac{1}{\gamma} \left[e^{\gamma(x - k_x a)} - 1 \right].$$

The result of Lemma 3 in the case $a_k = a$ for $k = 1, 2, \dots$ can be written as follows.

Lemma 4. *For every fixed $a > 0$ we have*

$$E\tau(x, \{a_k\}, \theta) = T(x) = \sum_{k=0}^{\infty} I_k(a) T_k(x, a).$$

The graphs of $T_k(x, a)$ and $T(x, a)$ are given on Fig. 3. The next properties one can derive, using the arguments of [2].

Lemma 5. The absolute minimum of $T_k(x, a)$ is

$$t_k = e^{\gamma[x + \ln \varphi(\gamma)] / (k+1)} \frac{1}{\gamma} \left[\frac{k}{\varphi(\gamma)} + 1 \right] - \frac{k+1}{\gamma}$$

and it gets in the point $\hat{a}_0 = x / (k+1) + \ln \varphi(\gamma) / \gamma(k+1) < x / (k+1)$.

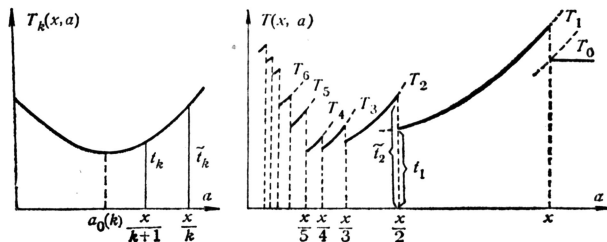


Fig. 3

Lemma 6. If $t_{k-1} = \lim_{a \downarrow x/k} T(x)$, $\tilde{t}_k = \lim_{t \uparrow x/k} T(x)$, then

$$t_{k-1} = \frac{k-1}{\gamma} \left(\frac{e^{\gamma x/k}}{\varphi(\gamma)} - 1 \right) + \frac{1}{\gamma} (e^{\gamma x/k} - 1), \quad \tilde{t}_k = \frac{k}{\gamma} \left(\frac{e^{\gamma x/k}}{\varphi(\gamma)} - 1 \right)$$

and always $t_{k-1} < \tilde{t}_k$.

Lemma 7. For any fixed x and γ

$$T_0(x) = \inf_a T(x) = \inf_k t_k = t_{k_0},$$

where k_0 is finite. The optimal copy interval is $a^* = x / (k_0 + 1)$.

Theorem 2. Under the assumptions of Theorem 1 for all X in the interval $[x_k, x_{k+1}]$ is fulfilled $a^* = X / (k+1)$, where $x_0 = 1$ and the sequence $\{x_k\}$ is determined for all k by the equation

$$(7) \quad \left[1 + \frac{k+1}{\varphi(\gamma)} \right] e^{\gamma x_k / (k+2)} - \left[1 + \frac{k}{\varphi(\gamma)} \right] e^{\gamma x_k / (k+1)} = 1.$$

Proof. Obviously, Lemma 7 gives one and the same number of copies k_0 for different x , because $k_0 = k_0(x)$ is a stepwise function of x . It follows that for any k there is a x_k , for which $k_0 = k$ or $k_0 = k+1$ and

$$(8) \quad T_0(x_k) = t_k = t_{k+1}.$$

Here t_k and $T_0(x)$ are as in Lemma 6 and Lemma 7. The equation (7) follows from (8) in view of the expressions for t_k , given in Lemma 6.

Corollary. In the case $P\{\theta = \text{const}\} = 1$ the thresholds x_k are solutions of the equations

$$e^{\gamma x_k / (k+2)} [1 + (k+1)e^{\gamma\theta}] - e^{\gamma x_k / (k+1)} [1 + ke^{\gamma\theta}] = 1.$$

Denoting $z = \gamma\theta$, $x_k = \theta y_k$, the last equation get over

$$(9) \quad e^{z y_k / (k+2)} [1 + (k+1)e^z] - e^{z y_k / (k+1)} [1 + ke^z] = 1.$$

Table 1

k	Thresholds $y_k, \theta=1$				
	$z^{-1}=30$	$z^{-1}=60$	$z^{-1}=90$	$z^{-1}=120$	$z^{-1}=120$
0	11.03	15.54	19.02	21.95	24.53
1	28.37	26.21	32.23	37.31	41.79
2	25.56	36.66	45.18	52.37	58.70
3	32.71	47.05	58.04	67.33	75.49
4	39.84	57.39	70.87	82.24	92.25
5	46.96	67.72	83.68	97.14	108.92
6	54.08	78.06	96.47	112.01	125.69
7	61.18	88.37	109.24	126.89	142.37
8	68.29	98.69	122.04	141.76	159.09
9	75.39	109.00	134.83	156.52	175.71
10	82.49	119.31	147.54	171.47	192.33
11	89.57	129.60	160.35	186.30	209.05
12	96.70	139.92	173.01	201.08	225.75
13	103.78	150.27	185.92	215.00	242.41
14	110.85	160.46	198.60	230.86	258.57
15	117.98	170.88	211.37	245.56	275.69
16	125.07	181.03	224.15	260.41	292.49
17	132.17	191.47	236.91	274.91	308.68
18	139.25	201.74	249.71	290.40	325.52
19	146.36	212.06	262.44	304.53	342.65

If $z = \gamma\theta = \text{const}$, then the thresholds satisfy the equality $x_k = \theta y_k$, where y_k are the solutions of (9). Hence, to have the thresholds x_k for any γ and θ it is enough to have the tables of the thresholds y_k , for $\theta = 1$ and for different values of z . The Table 1 gives us an example for the solutions of (9).

6.2. The case when X is r.v. and $P\{X < x\} = A(x)$. Let γ and the copy interval a be given, and $E e^{-s\theta} = \varphi(s)$. Denote $\tau(a) = E \tau(X, \{a\}, \theta)$.

Theorem 3. We have

$$\tau(a) = \frac{1}{\gamma} \sum_{k=0}^{\infty} e^{-\gamma ka} \int_{ka}^{(k+1)a} e^{\gamma x} dA(x) + \frac{1}{\gamma} \left[\frac{e^{\gamma a}}{\varphi(\gamma)} - 1 \right] \sum_{k=0}^{\infty} k [A((k+1)a) - A(ka)] - \frac{1}{\gamma}.$$

Proof. It follows from Lemma 4 and the equation

$$\tau(a) = \int_0^{\infty} E \tau(x, \{a\}, \theta) dA(x).$$

6.3. The case $P\{X > x\} = e^{-\lambda x}, \lambda > 0$. In this particular case the following statement holds.

Theorem 4. If

(i) $\lambda \neq \gamma$

then $\tau(a) = \frac{1}{\gamma} \left\{ \frac{\lambda}{\gamma - \lambda} [e^{(\gamma - \lambda)a} - 1] + \frac{e^{(\gamma - \lambda)a}}{\varphi(\gamma)} - 1 \right\} / (1 - e^{-\lambda a})$ and if

(ii) $\lambda = \gamma$

then $\tau(a) = \frac{1}{\gamma} [a\gamma + \frac{1}{\varphi(\gamma)} - 1] / [1 - e^{-\gamma a}]$.

Proof. The proof follows from Theorem 3. One must calculate the integrals and the obtained sums.

Theorem 5. *The optimal copy interval a^* is the unique solution of the equation*

$$(10) \quad [e^{-\gamma a} - e^{-\lambda a}]/(\gamma - \lambda) + (1 - e^{-\lambda a})/[\lambda \varphi(\gamma)] = [\varphi(\gamma) - 1]/[\gamma \varphi(\gamma)]$$

in the case $\lambda \neq \gamma$, and of the equation

$$(11) \quad (\gamma a - 1)e^{-\gamma a} = \varphi(\gamma)$$

in the case $\lambda = \gamma$.

If $\lambda \neq \gamma$, a^ exists if and only if*

$$(12) \quad \lambda < \gamma/[1 - \varphi(\gamma)].$$

Proof. The optimal copy interval a^* satisfies $\tau(a^*) = \inf_a \tau(a)$, and it is a solution of the equation

$$(13) \quad \frac{\partial}{\partial a} \tau(a) = 0.$$

The exact analysis of (13) for the functions $\tau(a)$ from Theorem 4 goes to (10) or (11). When the inequality (12) is not valid, the optimal copy interval is $a^* = \infty$.

7. Applications. For given γ and θ it is not difficult to calculate $a^* = a^*(\lambda)$ for the suitable values of λ ; a^* depends on γ and θ , which also can be taken into account. In the practice usually γ, θ are given for the process realization, but the pure time duration X can change (depending on λ). To make optimal copies one needs the tables of $a^*(\lambda)$.

8. Generalizations

a. To introduce a cost function $\alpha(x)$, instead of time duration $\tau(x)$.

If the process duration is x and there are not breakdowns, it costs $c(x)$; if the copy duration is θ , it costs $b(\theta)$; if the repeated time is y , it costs $h(y)$.

Under the assumptions: the breakdowns form a Poisson process with a parameter $\gamma > 0$, the copy duration θ is constant, the breakdowns may occur during a copy and the restart begins from the last successive copy, we conclude, that the copies must be made at equidistant moments of length a .

Theorem 6. *For a fixed time duration x of the process X (without breakdowns) and given a , the expected value of the cost function is*

$$\begin{aligned} A_k(x, a) &= E \alpha(x, a, \theta, \gamma) \\ &= e^{\gamma(a+\theta)} c(x) e^{-\gamma(x+k\theta)} + (e^{\gamma(a+\theta)} - 1) \sum_{v=1}^{k-1} c(x - va) e^{-\gamma(x - va + (k-v)a)} \\ &\quad + c(x - ka) + \int_0^{x-ka} (c(u) + h(u)) \gamma e^{-\gamma u} du \\ &\quad + k e^{\gamma(a+\theta)} \left(\int_0^a h(u) \gamma e^{-\gamma u} du + e^{-\gamma a} (1 - e^{-\gamma \theta}) h(a) + e^{-\gamma \theta} b(\theta) \right) \\ &\quad + \int_0^\theta b(u) \gamma e^{-\gamma u} du + \sum_{v=1}^k ((k-v+1) e^{\gamma(a+\theta)} - (k-v)) e^{-\gamma(v-1)\theta} \\ &\quad \times \left(\int_a^{va} c(u) \gamma e^{-\gamma u} du + e^{-\gamma va} (1 - e^{-\gamma \theta}) c(va) \right). \end{aligned}$$

The general form of the expected cost function is

$$A(x, a) = \sum_{k=0}^{\infty} I_k(a) \cdot A_k(x, a),$$

where $I_k(a)$ was introduced above.

Further one can use the same techniques to find the optimal value a_0 .

b. To use some more general forms of d.f. $A(x)$ and d.f. of $\{\xi_n\}$.

If $B(x) = P\{\xi_n < x\}$ is not the exponential distribution, then the copies will be not equidistant. We have an optimal scheduling problem, which can be solved if $B(x)$ has a monotone failure rate.

c. To introduce a process $\{X_n\}$ instead of one fixed random duration X .

d. To consider the repair time durations caused from catastrophes, which must be made before restarts.

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